Working across Time Zones: Exporters and the Gender Wage Gap

Esther Ann Bøler¹ Beata Javorcik² Karen Helene Ulltveit-Moe³

¹Imperial College Business School, ESOP, CEPR and CEP

²University of Oxford, CEPR and ESOP

³University of Oslo, CEPR and ESOP

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Motivation

- Globalization blamed for increasing inequality
 - What about the gender dimension?
- Women earn less than men even in Scandinavia
 - holds after controlling for worker characteristics, hours worked full vs. part time and occupation
- Exporters pay higher wages
 - holds after controlling for firm characteristics

Research question

Is the gender wage gap (GWG) higher at exporting firms?

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• Exporters may require a greater commitment from their employees

- Late night calls due to different time zones
- International travel & 24/7 availability
- ► Flexibility: e.g. greater responsiveness to unexpected problems
- Survey and empirical evidence on commitment being key to export success

Is the gender wage gap (GWG) higher at exporting firms?

• Exporters may require a greater commitment from their employees

- Late night calls due to different time zones
- International travel & 24/7 availability
- > Flexibility: e.g. greater responsiveness to unexpected problems
- Survey and empirical evidence on commitment being key to export success
- If women are less able to be flexible, or are perceived as such, they may be paid less

Survey evidence on perceived gender differences

- Employer surveys (e.g. from Iceland) reveal that
 - women are perceived as more family oriented than men
 - women are perceived as less committed and reliable than men
- A Survey by Slater & Gordon of managers in the UK (2014)
 - showed that a third of managers claim that women are not as good at their jobs when they come back from maternity leave
- Survey in Harvard Business Review
 - comparing persons with same qualification: mother deemed less competent than father

Are exporters more commitment intensive?

- Matched employer-employee data from Norway indicate that college educated workers are less likely to take sick leave if employed by an exporter
- True for both men and women

- The GWG is **lower** at exporting firms, but this result is **reversed** once unobservable worker-firm heterogeneity is controlled for
- In other words, women working for exporters are paid more than other women, but they are underpaid given their unobservable characteristics
- The GWG is systematically related to the **overlap in business hours** with the export markets

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Data 1996-2010

- Firm-level data from Statistics Norway's Capital database
 - unbalanced panel of all joint-stock companies in the Norwegian manufacturing sector
 - covers around 90% of manufacturing value added and employment
 - key variables: profits, employment, sector, location
- Employee register
 - key variables: wages, gender, age, children, education, occupational code (from 2003)
 - experience calculated from pensions register
 - full-time workers
- Firm-level customs trade data
- Centrality measure from Norwegian Institute for Urban and Regional Research (size of local labor market, proximity to urban centers)

Baseline model: Controlling for observables

- How large is the gender wage gap when we account for observable worker and firm characteristics?
- Is there a difference in the gender wage gap between exporters and non-exporters?
- A Mincer-type regression:

In
$$Wage_{ijst} = eta$$
 Fem $_i * Exporter_{jt} + X_{it}\lambda + \delta_{st} + arepsilon_{ijst}$

where i denotes worker, j firm, s sector, t year

• Worker characteristics: Experience, Experience squared, Education, Children and Centrality

Controlling for observable worker characteristics

	A (1)	All (2)	A (3)	A (4)	College (5)	College (6)	No College (7)	No College (8)
Female* Exporter	.058***	.054***	01**	008#	032***	030**	008	004
	(.007)	(.006)	(.005)	(.005)	(.010)	(.014)	(.005)	(.005)
Female	269***	264***						
	(.005)	(.004)						
Exporter	.082 ***	007 ***	.006 *		.011*		.006*	
	(.006)	(.003)	(.004)		(.006)		(.004)	
Education	.059***	.051***	.063***	.063***	.063***	.061***	.075***	.075***
	(.002)	(.001)	(.002)	(.002)	(800)	(.008)	(.003)	(.003)
Experience	.034***	.032***	.02***	.023***	.014	.021*	.041***	.037***
	(.000)	(.000)	(800.)	(.007)	(.011)	(.012)	(.010)	(.010)
Experience squared	058 ***	054***	055***	054***	066***	067***	051***	051***
	(.001)	(.001)	(.001)	(.001)	(.002)	(.002)	(.001)	(.001)
Child ren	026***	019***	066***	068***	035***	033***	083***	084***
	(.003)	(.002)	(.006)	(.006)	(.009)	(.010)	(.006)	(.006)
Centrality	.062***	.07***	.001	.01	.010	.011	003	006
	(.007)	(.012)	(.007)	(.006)	(.007)	(.010)	(.008)	(.005)
Industry-Year FE	Yes	Yes	Yes	No	Yes	No	Yes	No
Firm FE	No	Yes	No	No	No	No	No	No
Spell FE	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	No	No	No	Yes	No	Yes	No	Yes
0								
	2,113,023	2,113,023	2,113,023	2,113,623	110,800	110,666	2,100,012	2,100,012
K-squared	.40	.47	.75	.83	.81	.89	.70	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. ***= p-val<.01, ** = p-val<.05, * = p-val<.11 # = p-val<.11

- Working for an exporting firm seems to close the gender wage gap by about 6% points (or a fifth)
- Holds even when we allow all the coefficients to vary by export status and gender

But what about differences across firms?

- Control for firm fixed effects
- Identification comes from firms changing their export status
- A Mincer-type regression:

 $In Wage_{ijst} = \beta Fem_i * Exporter_{jt} + X_{it}\lambda + \delta_{st} + \delta_j + \varepsilon_{ijst}$

where i denotes worker, j firm, s sector, t year

 Worker characteristics: Experience, Experience squared, Education, Children and Centrality

. . . and unobservable firm heterogeneity

	All (1)	All (2)	All (3)	All (4)	College (5)	College (6)	No College (7)	No College (8)
Female* Exporter	.058***	.05 4***	01**	008#	032***	030**	008	004
Female	(.007) - <mark>.269</mark> ***	(.006) <mark>264</mark> ***	(.005)	(.005)	(.010)	(.014)	(.005)	(.005)
European	(.005) 082 ***	(.004) 007 ***	006 *		0.1.1*		00.6*	
Exporter	(.006)	(.003)	(.004)		(.006)		(.004)	
Education	.059***	.05 1***	.063***	.063***	.063***	.061***	.075***	.075***
	(.002)	(.001)	(.002)	(.002)	(.008)	(800.)	(.003)	(.003)
Experience	.034***	.032***	.02***	.023***	.014	.021*	.041***	.037***
	(.000)	(.000)	(800.)	(.007)	(.011)	(.012)	(.010)	(.010)
Experience squared	058***	054***	055***	054***	066***	067***	051***	051***
	(.001)	(.001)	(.001)	(.001)	(.002)	(.002)	(.001)	(.001)
Child ren	026 ***	019***	066***	068***	035***	033***	083***	084***
	(.003)	(.002)	(.006)	(.006)	(.009)	(.010)	(.006)	(.006)
Centrality	.062***	.07***	.001	.01	.010	.011	003	006
	(.007)	(.012)	(.007)	(.006)	(.007)	(.010)	(.008)	(.005)
Industry-Year FE	Yes	Yes	Yes	No	Yes	No	Yes	No
Firm FE	No	Yes	No	No	No	No	No	No
Spell FE	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	No	No	No	Yes	No	Yes	No	Yes
Obs	2,713,623	2,713,623	2,713,623	2,713,623	553,611	553,611	2,160,012	2,160,012
R-squared	.40	.47	.75	.83	.81	.89	.70	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. ***= p-val<.01, ** = p-val<.05, * = p-val<.1,# = p-val<.11

But what about differences in worker ability, etc.?

- Control for worker-firm (spell) fixed effects
- Identification comes from workers employed in the firm as the firm changes its export status
- A Mincer-type regression:

$$In Wage_{ijst} = \beta Fem_i * Exporter_{jt} + X_{it}\lambda + \delta_{st} + \delta_{ij} + \varepsilon_{ijst}$$

where i denotes worker, j firm, s sector, t year

• Worker characteristics: Experience, Experience squared, Education, Children and Centrality

. . . and unobservable worker-firm heterogeneity

	A	All	A	All	College	College	No College	No College
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Female* Exporter	.058***	.05 4***	01**	008#	- 032***	030**	008	004
Female	(.007) 269***	(.006) <mark>264</mark> ***	(.005)	(.005)	(.010)	(.014)	(.005)	(.005)
Exporter	(.005) .082 *** (.005)	(.004) 007 *** (.003)	.006 *		.011*		.006*	
Education	.059****	.051***	.063***	.063***	.063***	.061***	.075***	.075***
	(.002)	(.001)	(.002)	(.002)	(.008)	(.008)	(.003)	(.003)
Experience	.034***	.032***	.02***	.023***	.014	.021*	.041***	037***
	(.000)	(.000)	(.008)	(.007)	(.011)	(.012)	(.010)	(010)
Experience squared	058***	054***	055***	054***	066***	067***	051****	051***
	(.001)	(.001)	(.001)	(.001)	(.002)	(.002)	(.001)	(.001)
Child ren	026***	019***	066***	068***	035****	033****	083****	084***
	(.003)	(.002)	(.006)	(.006)	(.009)	(.010)	(.006)	(.006)
Centrali ty	.062****	.07***	.001	.01	.010	.011	003	006
	(.007)	(.012)	(.007)	(.006)	(.007)	(.010)	(.008)	(.005)
Industry-Year FE	Yes	Yes	Yes	No	Yes	No	Yes	No
Firm FE	No	Yes	No	No	No	No	No	No
Spell FE	No	No	<mark>Yes</mark>	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	No	No	No	Yes	No	Yes	No	Yes
Obs	2,713,623	2,713,623	2,713,623	2,713,623	553,611	553,611	2,160,012	2,160,012
R-squared	.40	.47	.75	.83	.81	.89	.70	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. ***= p-val<.01, ** = p-val<.05, * = p-val<.1,# = p-val<.11

But what if entry into exporting is not exogenous?

- Control for worker-firm (spell) as well as firm-year fixed effects
- Identification comes from workers employed in the firm as the firm changes its export status
- A Mincer-type regression:

$$ln Wage_{ijst} = \beta Fem_i * Exporter_{jt} + X_{it}\lambda + \delta_{ij} + \delta_{jt} + \varepsilon_{ijst}$$

where i denotes worker, j firm, s sector, t year

• Worker characteristics: Experience, Experience squared, Education, Children and Centrality

. . . and firm-year unobservables

	A (1)	All (2)	A II (3)	All (4)	College (5)	College (6)	No College (7)	No College (8)
Female* Exporter	.058***	.054***	01**	008#	032***	030**	008	004
Female	(.007) - .269***	(.006) 264 ***	(.005)	(.005)	(.010)	(.014)	(.005)	(.005)
	(.005)	(.004)						
Exporter	.082 ***	007 ***	.006 *		.011*		.006*	
Education	(.006) .059***	(.003) .051***	(.004) .063***	.063***	(.006) .063***	.06 1***	(.004) .075***	.075***
	(.002)	(.001)	(.002)	(.002)	(.008)	(800.)	(.003)	(.003)
Experience	.034	.032	.02	.023	.014	.021*	.041	.037
Experience squared	058***	054***	055***	054***	066***	067***	051***	051***
Children	(.001) - 026***	(.001) - 019***	(.001) - 066***	(.001) - 068***	(.002) - 035***	(.002) - 033***	(.001) - 083***	(.001) - 084***
ennaren -	(.003)	(.002)	(.006)	(.006)	(.009)	(.010)	(.006)	(.006)
Centrality	.062***	.07***	.001	.01	.010	.011	003	006
	(.007)	(.012)	(.007)	(.006)	(.007)	(.010)	(.008)	(.005)
Industry-Year FE Firm FE	<mark>Yes</mark> No	Yes Yes	Yes No	No No	Yes No	No No	Yes No	No No
Spell FE Firm-Year FE	No No	No No	<mark>Yes</mark> No	Yes <mark>Yes</mark>	Yes No	Yes Yes	Yes No	Yes Yes
Obs R annual	2,713,623	2,713,623	2,713,623	2,713,623	553,611	553,611	2,160,012	2,160,012
n-squareu	.+0	. + 7	.75	.03	.01	.09	.70	.00

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. ***= p-val<.01, ** = p-val<.05, * = p-val<.1,# = p-val<.11

• Exporters may require a greater commitment from their employees

- Late night calls due to different time zones
- International travel & 24/7 availability .
- Flexibility: e.g. greater responsiveness to unexpected problems
- This is more likely to apply to skilled workers who negotiate deals with clients, provide technical advice and support, take care of logistics such as transport and customs clearance, etc.

College educated vs. other workers

	A (1)	All (2)	A (3)	All (4)	College	College	No College	No College (8)
Female* Exporter	.058***	.05 4***	01**	.008#	032***	- 030**	008	.004
Female	(.007) - 269 ***	(006) - 264 ***	(.005)	(.005)	(.010)	(.014)	(.005)	(.005)
	(.005)	(.004)						
Exporter	.082 ***	007 ***	.006 *		.011*		.006*	
	(.006)	(.003)	(.004)		(.006)		(.004)	
Education	.059***	.05 1***	.063***	.063***	.063***	.061***	.075***	.075***
	(.002)	(.001)	(.002)	(.002)	(.008)	(.008)	(.003)	(.003)
Experience	.034***	.032***	.02***	.023 ***	.014	.021*	.041***	.037***
	(.000)	(.000)	(.008)	(.007)	(.011)	(.012)	(.010)	(.010)
Experience squared	058***	054***	055***	054***	066***	067***	051***	051***
	(.001)	(.001)	(.001)	(.001)	(.002)	(.002)	(.001)	(.001)
Child ren	026 ***	019***	066***	068***	035***	033***	083***	084 ***
	(.003)	(.002)	(.006)	(.006)	(.009)	(.010)	(.006)	(.006)
Centrality	.062***	.07***	.001	.01	.010	.011	003	006
	(.007)	(.012)	(.007)	(.006)	(.007)	(.010)	(800.)	(.005)
Industry-Year FE	Yes	Yes	Yes	No	Yes	No	Yes	No
Firm FÉ	No	Yes	No	No	No	No	No	No
Spell FE	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	No	No	No	Yes	No	Yes	No	Yes
Obs	2,713,623	2,713,623	2,713,623	2,713,623	553,611	553,611	2,160,012	2,160,012
R-squared	.40	.47	.75	.83	.81	.89	.70	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. ***= p-val<.01, ** = p-val<.05, * = p-val<.1

- Customs data give us information on the firm's exports by 8-digit HS product and destination country
- We proxy for heterogeneity among exporters using
 - Export share: ratio of exports to total production
 - > Number of destinations to which a firm exports in a given year
 - Number of exported varieties: the number of 8-digit HS product-destination country combinations

Heterogeneity among exporters

	Mean	Median	Min	Max
Export share	0.18	0.04	<0.01	1
No of destinations	7.31	3	1	146
No of varieties	24.88	6	1	4,126

Table: Exporter descriptives

Notes: All numbers are based on the panel of firm-level data for 1996-2010. An exporter has by definition exports abo NOK 10,000 (USD 1,100). A variety is defined as product-destination combination. <0.01 denotes export share below percent.

Heterogeneity among exporters

	College (1)	College (2)	College (3)	No College (4)	No College (5)	No College (6)
Female*Export share	014** (.007)			011 (.011)		
Female*Destinations		013** (.006)			007* (.003)	
Female*No of varieties			009** (.004)			006** (.002)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	454,063	494,601	494,601	1,620,597	1,712,387	1,712,387
R-squared	.88	.88	.88	.79	.80	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls inclued education, experience, experience squared, children and centrality. *** = p-va| < .01, ** = p-va| < .05, * = p-va| < .1

- An 10% point increase in the export share is associated with the GWG going up by 13% points for college educated workers
- Adding 4 additional export markets leads to the GWG going up by 5% points
- Adding 10 additional export varieties implies an 8% point increase in the GWG

- Compile data on time zone(s) for each export destination country
- For countries spanning multiple time zones, consider only mainland and use the average value
- Calculate the the overlap in business hours between each destination country and Norway assuming the standard 9am to 5pm office hours
- Assume that the need for communications with a particular country increases in the number of traded products, take the average business hour overlap for each firm across country-product combinations in each year

Table: Exporter descriptives

	Mean	Median	Min	Max
Export share	0.18	0.04	<0.01	1
Business hour overlap	7.01	7.67	0	8
No of destinations	7.31	3	1	146
No of varieties	24.88	6	1	4,126

Notes: All numbers are based on the panel of firm-level data for 1996-2010. An exporter has by definition exports abo NOK 10,000 (USD 1,100). A variety is defined as product-destination combination. <0.01 denotes export share below percent.

				All Workers		
	College	College	College	No College	No College	No College
	(1)	(2)	(2)	(3)	(4)	(4)
Female*/n Business hours	.028*		.033**	.001		.001
	(.015)		(.015)	(.009)		(.009)
Female* <i>In</i> Gender index		014	082		.006	.012
		(.099)	(.102)		(.048)	(.052)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	493,736	494,601	493,177	1,708,945	1,709,121	1,705,848
R-squared	.88	.88	.88	.80	.80	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls inclueducation, experience, experience squared, children and centrality. *** = p-val<.01, ** = p-val<.05, * = p-val<.1 • A 4 hour decrease in the business hour overlap is associated with the GWG increasing by 11% points for college educated women.

- But could our result be driven by discrimation against women in the Middle East and Asia?
- Perhaps women are less effective at doing business there and this explains the larger GWG in firms exporting to far away countries?
- To address this issue we control for the **Gender Gap Index** from the World Economic Forum
 - The index pertains to 2006 and covers 115 countries
 - The index measures gaps between men and women, rather than absolute levels of female outcomes
 - It focuses on outcomes rather than input measures.
 - It rewards countries for closing the gender gap, but neither rewards not penalizes countries where women have an advantage over men in certain outcomes
 - As with the Business hour overlap, we take the average of the index over all product-destination country combinations observed in a particular year

				All Workers		
	College	College	College	No College	No College	No College
	(1)	(2)	(2)	(3)	(4)	(4)
Female* <i>In</i> Business hours	.028*		.033**	.001		.001
	(.015)		(.015)	(.009)		(.009)
Female* <i>In</i> Gender index		014	082		.006	.012
		(.099)	(.102)		(.048)	(.052)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	493,736	494,601	493,177	1,708,945	1,709,121	1,705,848
R-squared	.88	.88	.88	.80	.80	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls inclueducation, experience, experience squared, children and centrality. *** = p-val<.01, ** = p-val<.05, * = p-val<.1

				All Workers		
	College	College	College	No College	No College	No College
	(1)	(2)	(2)	(3)	(4)	(4)
Female* <i>In</i> Business hours	.028*		.033**	.001		.001
	(.015)		(.015)	(.009)		(.009)
Female* <i>In</i> Gender index		014	082		.006	.012
		(.099)	(.102)		(.048)	(.052)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	493,736	494,601	493,177	1,708,945	1,709,121	1,705,848
R-squared	.88	.88	.88	.80	.80	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls inclueducation, experience, experience squared, children and centrality. *** = p-val<.01, ** = p-val<.05, * = p-val<.1

- Women with young children are likely to be less flexible
- The decision to have children and its timing may be endogenous, so we focus on workers in their reproductive years
- We expect exporting to matter more for the GWG among the group of younger workers

	Workers under 45 years of age							
	College (1)	College (2)	College (3)	No College (4)	No College (5)	No College (6)		
Female*/n Business hours overlap	.042**		.049**	.005		036		
	(.020)		(.021)	(.012)		(.080)		
Female*/n Gender index		027	.129		030	036		
		(.124)	(.133)		(.074)	(.013)		
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes		
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes		
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes		
Obs	327,973	328,142	327,559	1,028,554	1,028,674	1,026,584		
R-squared	.88	.88	.88	.81	.81	.81		

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered by firm. Worker controls inclued ucation, experience, experience squared, children and centrality. *** = p-va| < .01, ** = p-va| < .05, * = p-va| < .1

- Dummy for one child versus more children
- Controlling for occupational fixed effects
- Allowing the GWG to differ by occupation

Robustness checks

	College (1)	College (2)	College (2)	All Workers No College (3)	No College (4)	No College (4)
Female* <i>In</i> Business hours overlap	.033**	.039**	.039**	.001	001	002
	(.015)	(.015)	(.015)	(.009)	(.017)	(.017)
Female*/n Gender index	082	128	128	.014	.027	.028
	(.101)	(.134)	(.134)	(.053)	(.070)	(.070)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Controlling for 1 versus 2 children	Yes	No	No	Yes	No	No
Occupation FE	No	Yes	Yes	No	Yes	Yes
Occupation*Female FE	No	No	Yes	No	No	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	493,177	288,444	288,444	1,705,848	888,713	888,713
R-squared	.88	.89	.89	.80	.82	.82

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls inclued education, experience, experience squared, children and centrality. *** = p-val<.01, ** = p-val<.05, * = p-val<.1

Robustness checks

	College (1)	College (2)	College (2)	All Workers No College (3)	No College (4)	No College (4)
Female*/n Business hours overlap	.033**	.039**	.039**	.001	001	002
	(.015)	(.015)	(.015)	(.009)	(.017)	(.017)
Female*/n Gender index	082	128	.128	.014	.027	.028
	(.101)	(.134)	(.134)	(.053)	(.070)	(.070)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Controlling for 1 versus 2 children	Yes	No	No	Yes	No	No
Occupation FE	No	Yes	Yes	No	Yes	Yes
Occupation*Female FE	No	No	Yes	No	No	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	493,177	288,444	288,444	1,705,848	888,713	888,713
R-squared	.88	.89	.89	.80	.82	.82

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls inclued education, experience, experience squared, children and centrality. *** = p-val<.01, ** = p-val<.05, * = p-val<.1

Alternative explanation: Taste-based discrimination

- All firms would like to engage in taste-based discrimination, but only exporters are protable enough to do so
- Only most productive exporters sell to far away destinations
- Allow the GWG to differ with profitability, firm size and multinational status

Alternative explanation: Taste-based discrimination

	All workers					
	College	College	College	No College	No College	No College
	(1)	(2)	(2)	(3)	(4)	(4)
Female*In Business hours overlap	.028*		.034**	.001		.001
	(.015)		(.015)	(.009)		(.009)
Female* <i>In</i> Gender index		014	082		.006	.012
		(.099)	(.102)		(.048)	(.053)
Female*Profitability	000	000	000	002**	002**	002**
	(.001)	(.001)	(.001)	(.005)	(.001)	(.001)
Female*MNC	002	002	002	.002	.002	.001
	(.008)	(.008)	(.008)	(.006)	(.006)	(.006)
Female*/n Size	004	003	004	008	008	008
	(.008)	(.008)	(.008)	(.005)	(.005)	(.005)
Worker Controls	Yes	Yes	Yes	Yes	Yes	Yes
Spell FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Obs	493,736	494,016	493,177	1,708,945	1,709,121	1,705,848
R-squared	.88	.88	.88	.80	.80	.80

Notes: Dependent variable is log wage. *MNC* is a dummy that equals one if the firm is registered as having a posit ownership share in a firm located in a foreign country in the year of observation, and zero otherwise. Estimates are bas on the panel of worker-level data for 1996-2010. Standard errors in parentheses clustered on firm. Worker controls inclu education, experience, experience squared, children and centrality. ***= p-val<.01, ** = p-val<.05, * = p-val<.1

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- Is there something special about college workers?
- Not really, splitting the sample into Managers & Professionals & Technicians vs. Other Occupations leads to the same conclusions
- Splitting occupations into those focused on Cognitive task vs Routine tasks does not
- Supports the flexibility story

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Alternative sample split

	All Workers Managers & Professionals & Technicians (1)	Other Occupations (2)
Female*/n Business hours overlap	.032*	.002
	(.018)	(.016)
Female* <i>In</i> Gender index	011	025
	(.096)	(.079)
Worker Controls	Yes	Yes
Spell FE	Yes	Yes
Firm-Year FE	Yes	Yes
Obs	353,725	791,022
R-squared	.89	.80

Notes: Dependent variable is log wage. Estimates are based on the panel of worker-level data for 1996-2010. O workers employed by exporting firms are included. Standard errors in parentheses clustered on firm. Worker controls inclu education, experience, experience squared, children and centrality. ***= p-val<.01, ** = p-val<.05, * = p-val<.1

Conclusions

- Globalization matters for gender outcomes
- The GWG is **lower** at exporting firms, but this result is **reversed** once unobservable worker-firm heterogeneity is controlled for
- Put differently, women working for exporters are paid more than other women, but they are underpaid given their unobservable characteristics
- There is heterogeneity within the group of exporters: GWG is systematically related to the **overlap in business hours** with the export markets

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